

Medical image analysis using high-dimensional information-theoretic criteria

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Motivation

- ▶ Information theory provides powerful criteria for solving low-level problems in medical image understanding based on **statistical descriptions of visual features**
 - Registration & Motion estimation
 - frame-to-frame similarity → **informations**
 - Segmentation & Tracking
 - object / background discrepancy → **divergences**
 - frame / labelling similarity → **informations**
 - Classification & Retrieval
 - Class separation → **divergences**
 - Class consistency → **informations**

Motivation

- ▶ In medical imaging: a huge, still expanding literature body on mutual information and its variations since 1995

Tutorial

MICCAI 2009 tutorial

Information theoretic similarity measures for image registration and segmentation

ubimon.doc.ic.ac.uk/MICCAI09/a1882.html

Review article

J.P.W. Pluim, J.B.A. Maintz and M.A. Viergever

Mutual-information-based registration of medical images: a survey

IEEE Transactions on Medical Imaging, 22(7):986-1004, July 2003

Motivation

► Impact on the medical imaging domain

- In 2000, recognized by IEEE as *a landmark in the profession, with enduring importance and influence far beyond its peer*
- In 2005, princeps papers recognized by ISI as *one of the 10 most cited papers of the last decade published in Engineering* [Collignon *et al.*, 1995] [Viola and Wells, 1995]

Motivation

► Software implementations

- **AIR** (UCLA, Loni): www.loni.ucla.edu/Software/AIR
- **DROP** (TU Muenchen): www.mrf-registration.net
- **Elastix** (ISI, Utrecht): elastix.isi.uu.nl
- **FSL/FLIRT/FNIRT** (FMRIB, Oxford): www.fmrib.ox.ac.uk/fsl
- **ImageFusionTM**:
www.radionics.com/products/functional/imagefusion.shtml
- **IRTK** (Imperial College, London):
www.doc.ic.ac.uk/~dr/software
- **Slicer** (BWH): www.slicer.org
- **SPM** (University College, London): fil.ion.ucl.ac.uk/spm
- ...

Motivation

- ▶ In computer vision: growing interest for information-theoretic region-based segmentation via curve evolution since 2001
 - Statistical region competition paradigm
[Zhu and Yuille *et al.*, 1996]
 - Levelset framework
 - Optimization
 - standard variational calculus
[Freedman, 2004] [Kim *et al.*, 2005] [Rougon *et al.*, 2006]
[Michailovich *et al.*, 2007]
 - shape calculus
[Aubert *et al.*, 2003] [Jehan-Besson *et al.*, 2003]
[Herbulot *et al.*, 2006] [Boltz *et al.*, 2008]

Motivation

► Strong assets

- Solid theoretical foundations
- Blindness: no limitations on sensors / data
 - In particular: deal with multimodal / uncommensurable data
- Robustness against image degradations
(noise, artifacts, local distortions)
- No supervision
- “Easy” implementation
- Versatility and pervasiveness

Motivation

► BUT

- Notoriously difficult to reliably estimate in **high-dimensions**
- No straightforward generalization to the **multivariate** case

Restriction to bivariate information-theoretic measures acting on low-dimensional (usually scalar) random variables

► A stringent restriction

- Local intensity information only is ambiguous
 - **spatial** information (*i.e.* geometry) required
- Non-local, complex image features can be informative
 - *e.g.* wavelet coefficients, image tags, nonlocal textural features, etc.
- Multiple image analysis ?

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- ▶ **A. Collignon, F. Maes, D. Delaere, D. Vandermeulen, P. Suetens and G. Marchal**
Automated multimodality image registration using information theory
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Active contours for tracking distributions
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Region competition: Unifying snakes, region growing and Bayes/MDL for multiband image segmentation
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Outline

- 1 Foundations
- 2 kNN entropy estimators
- 3 Entropic graphs
- 4 Multivariate information measures

Outline

- 1 Foundations
- 2 kNN entropy estimators
- 3 Entropic graphs
- 4 Multivariate information measures

Notations

X (continuous/discrete) random variable (RV) over some d -dimensional state space \mathcal{X} with density p^X

$X(\Omega)$ set of i.i.d. samples of X , denoted $X(s)$ or X_s , indexed by $s \in \Omega$

Ω the image domain/grid, an image region or a point distribution

$X|Y$ conditional RV with density $p^{X|Y}$

(X, Y) joint RV with density $p^{X,Y}$

Differential entropy

Differential entropy

$$H(X) = - \int_{\mathcal{X}} p^X(x) \log p^X(x) dx$$

- Discrete framework: [Shannon entropy](#). Does **not** converge to differential entropy in the continuous limit

► Joint entropy

$$H(X, Y) = - \int_{\mathcal{X}^2} p^{X,Y}(x, y) \log p^{X,Y}(x, y) dxdy$$

► Conditional entropy

$$H(X|Y) = -E_Y \left(\int_{\mathcal{X}} p^{X|Y}(x|y) \log p^{X|Y}(x|y) dx \right)$$

Differential entropy estimation

Notice that:

$$H(X) = -\mathbb{E}_X [\log p^X]$$

Ahmad-Lin entropy estimator [Ahmad and Lin, 1976]

If p^X is known

$$H^{AL}(X) = -\frac{1}{|\Omega|} \sum_{s \in \Omega} \log p^X(X_s)$$

is a **consistent** estimator of differential entropy $H(X)$

- Holds if p^X is replaced by a **consistent** estimator \hat{p}^X of p^X
- **Ergodic approximation:** expectation over the state space \mathcal{X} is replaced by averaging over the sample set $X(\Omega)$
- **Levelset segmentation:** Ω is a free boundary domain and p^X is domain-dependent → **shape calculus** required

Kullback-Leibler divergence

- Discrepancy of a RV X w.r.t. a reference RV Y

Kullback-Leibler divergence

$$D_{\text{KL}}(X \parallel Y) = \int_X p^X(x) \log \frac{p^X(x)}{p^Y(x)} dx$$

- Non-symmetric
- Interpretation as an entropy w.r.t. the measure $p^Y dx$

- Symmetrized KL divergence

$$D_{\text{KL}}(X, Y) = \frac{1}{2} \left[D_{\text{KL}}(X \parallel Y) + D_{\text{KL}}(Y \parallel X) \right]$$

Kullback-Leibler divergence

- Discrepancy of a RV X w.r.t. a reference RV Y

Kullback-Leibler divergence

$$D_{\text{KL}}(X \parallel Y) = \int_X \frac{p^X(x)}{p^Y(x)} \log \frac{p^X(x)}{p^Y(x)} p^Y(x) dx$$

- Non-symmetric
- Interpretation as an entropy w.r.t. the measure $p^Y dx$

- Symmetrized KL divergence

$$D_{\text{KL}}(X, Y) = \frac{1}{2} \left[D_{\text{KL}}(X \parallel Y) + D_{\text{KL}}(Y \parallel X) \right]$$

Kullback-Leibler divergence estimation

Notice that:

$$D_{\text{KL}}(X \parallel Y) = \mathbb{E}_X \left[\log \frac{p^X}{p^Y} \right]$$

Ahmad-Lin KL divergence estimator

If p^X and p^Y are known,

$$D_{\text{KL}}^{\text{AL}}(X \parallel Y) = -\frac{1}{|\Omega|} \sum_{s \in \Omega} \log \frac{p^X(X_s)}{p^Y(X_s)}$$

is a **consistent** estimator of KL divergence $D_{\text{KL}}(X \parallel Y)$.

- Holds if p^X, p^Y are replaced by **consistent** estimators \hat{p}^X, \hat{p}^Y

Mutual information

- ▶ Similarity between by 2 jointly observed RVs

Mutual information

Information shared by 2 jointly observed RVs

$$I(X, Y) = H(X) + H(Y) - H(X, Y)$$

- Reduction of uncertainty on a RV brought by another one

$$\begin{aligned} I(X, Y) &= H(X) - H(X|Y) \\ &= H(Y) - H(Y|X) \end{aligned}$$

- Deviation from independence

$$\begin{aligned} I(X, Y) &= \int_{\mathcal{X}^2} p^{X,Y}(x,y) \log \frac{p^{X,Y}(x,y)}{p^X(x)p^Y(y)} dx dy \\ &= D_{KL}((X, Y) \parallel X \times Y) \end{aligned}$$

Normalized information measures

Normalized Mutual Information [Studholme *et al.*, 1999]

$$NMI(X, Y) = \frac{H(X) + H(Y)}{H(X, Y)} = 1 + \frac{I(X, Y)}{H(X, Y)}$$

Robust w.r.t. incomplete statistics

- Registration: partial overlap of image supports

Entropy Correlation Coefficient [Maes *et al.*, 1997]

$$ECC(X, Y) = \frac{2I(X, Y)}{H(X) + H(Y)} = 2 \left(1 - \frac{1}{NMI(X, Y)} \right)$$

Also called Symmetric Uncertainty Coefficient
[Melbourne *et al.*, 2009]

Exclusive information

Exclusive information [Rougon *et al.*, 2003] [Zhang *et al.*, 2005]

Information contained solely in either of 2 jointly observed RVs

$$\begin{aligned} Z(X, Y) &= I(X, Y) - H(X, Y) \\ &= H(X) + H(Y) - 2H(X, Y) \\ &= -[H(X|Y) + H(Y|X)] \end{aligned}$$

- Faster optimization than MI

Shannon information theory

Entropy $H(X)$

$$-\mathbb{E}_X [f(p^X)]$$



Divergence $D(X \parallel Y)$

$$-H(X \parallel Y)$$



Mutual information $I(X, Y)$

$$D((X, Y) \parallel X \times Y)$$

- The **information metrics** f is the Kullback metrics $f(x) = \log(x)$
- This design can be generalized by considering other metrics f
→ **Ali-Silvey class**

f -entropy of a probability measure

μ, ν : probability measures (PMs). ν chosen as a **reference**.

Integral f -entropy

$$H_{f,\nu}(\mu) = - \int f\left(\frac{d\mu}{d\nu}\right) d\nu$$

f continuous, convex over \mathbb{R}^+

Non-integral f -entropy

$$H_{\psi,\nu}(\mu) = - \log \psi^{-1} \left(\int \frac{d\mu}{d\nu} \psi \left(\frac{d\mu}{d\nu} \right) d\nu \right)$$

ψ continuous, monotonic over \mathbb{R}^+

f, ψ : information metrics

f -entropy of a RV

- RVs are a special case: ν Borel ($d\nu = dx$) and $d\mu = p^X(x)d\nu$

Integral f -entropy

$$H_f(X) = - \int f(p^X(x)) dx$$

f continuous, convex over \mathbb{R}^+

Non-integral f -entropy

$$H_{\psi,\nu}(\mu) = - \log \psi^{-1} \left(\int p^X(x) \psi(p^X(x)) dx \right)$$

ψ continuous, monotonic over \mathbb{R}^+

f, ψ : information metrics.

Some integral f -entropies

- $f(x) = x \log x \rightarrow$ Differential entropy
- $f(x) = \frac{x-x^\alpha}{\alpha-1} \quad (\alpha \neq 1)$

Havrda-Charvát (Tsallis) entropy

$$H_\alpha(X) = \frac{1}{\alpha-1} \left(1 - \int [p^X(x)]^\alpha dx \right)$$

Properties:

- $\lim_{\alpha \rightarrow 1} H_\alpha(X) = H(X)$
- **Non-additivity:** if X, Y independent

$$H_\alpha(X, Y) = H_\alpha(X) + H_\alpha(Y) + (1 - \alpha)H_\alpha(X)H_\alpha(Y)$$

Some non-integral f -entropies

- $\psi(x) = \log x \rightarrow$ Differential entropy

The *only* both integral and non-integral f -entropy is $H(X)$

- $\psi(x) = x^{r-1} \quad (r \neq 1)$

Renyi entropy

$$H_r(X) = \frac{1}{1-r} \log \int [p^X(x)]^r dx$$

$r = \frac{1}{2} \rightarrow$ Bhattacharyaa entropy

Properties:

- $\lim_{r \rightarrow 1} H_r(X) = H(X)$
- **Additivity:** if X, Y independent:

$$H_r(X, Y) = H_r(X) + H_r(Y)$$

f -divergence of PMs

Integral / Non-integral f -divergence

$$D_f(\mu \parallel \nu) = -H_{f,\nu}(\mu)$$

f -divergence of RVs

- RVs are a special case: $d\mu = p^X(x)dx$ and $d\nu = p^Y(x)dx$

Integral f -divergence

$$D_f(X \parallel Y) = \int p^Y(x)f\left(\frac{p^X(x)}{p^Y(x)}\right) dx$$

Non-integral f -divergence

$$D_\psi(X \parallel Y) = \log \psi^{-1} \left(\int p^X(x)\psi\left(\frac{p^X(x)}{p^Y(x)}\right) dx \right)$$

- $D_f(X, Y) = \mathbb{E}_Y \left[f\left(\frac{p^X}{p^Y}\right) \right]$
- $D_\psi(X, Y) = \mathbb{E}_X \left[\psi\left(\frac{p^X}{p^Y}\right) \right]$

→ Ahmad-Lin-like estimators

Some f -divergences

- $f(x) = x \log x$ and $\psi(x) = \log x \rightarrow$ KL divergence

- $f(x) = \frac{x-x^\alpha}{\alpha-1} \quad (\alpha \neq 1)$
- $\psi(x) = x^{r-1} \quad (r \neq 1)$

I_α (Chernoff) divergence

$$D_\alpha(X \parallel Y) = \frac{1}{\alpha-1} \left(\mathcal{I}_q^{X \parallel Y} - 1 \right)$$

Renyi divergence

$$D_r(X \parallel Y) = \frac{1}{r-1} \log \mathcal{I}_q^{X \parallel Y}$$

where:

$$\mathcal{I}_q^{X \parallel Y} = \int_{\Omega} [p^X(x)]^q [p^Y(x)]^{1-q} dx$$

f-information of PMs

Integral / Non-integral *f*-information

$$I_f(\mu, \nu) = D_f((\mu, \nu) \parallel \mu \times \nu)$$

f -information of RVs

Integral f -information

$$I_f(X, Y) = \int p^X(x)p^Y(y) f\left(\frac{p^{X,Y}(x,y)}{p^X(x)p^Y(y)}\right) dx dy$$

Non-integral f -information

$$I_\psi(X, Y) = \log \psi^{-1} \left(\int p^{X,Y}(x,y) \psi\left(\frac{p^{X,Y}(x,y)}{p^X(x)p^Y(y)}\right) dx dy \right)$$

Note that $\frac{p^{X,Y}}{p^X p^Y}$ is the **likelihood ratio** between (X, Y) and $X \times Y$

- $I_f(X, Y) = \mathbb{E}_{X \times Y} \left[f\left(\frac{p^{X,Y}}{p^X p^Y}\right) \right]$ → Ahmad-Lin-like estimators
- $I_\psi(X, Y) \sim \mathbb{E}_{(X, Y)} \left[\psi\left(\frac{p^{X,Y}}{p^X p^Y}\right) \right]$

Some f -informations

- $f(x) = x \log x$ and $\psi(x) = \log x \rightarrow$ Mutual information
- $f(x) = \frac{x-x^\alpha}{\alpha-1}$ ($\alpha \neq 1$)
- $\psi(x) = x^{r-1}$ ($r \neq 1$)

I_α information

$$I_\alpha(X, Y) = \frac{1}{\alpha-1} \left(\mathcal{I}_\alpha^{X,Y} - 1 \right)$$

Renyi information

$$I_r(X, Y) = \frac{1}{r-1} \log \mathcal{I}_r^{X,Y}$$

where:

$$\mathcal{I}_q^{X,Y} = \int [p^{X,Y}(x,y)]^q [p^X(x)p^Y(y)]^{1-q} dx dy$$

Non-parametric density estimation

- Given a sample set $X(\Omega)$, the density $p^X(x)$ is the proportion $\frac{k(x)}{|\Omega|}$ of samples x per volume unit
 - Consider open balls $\mathcal{B}_r(x)$ of radius r centered at x
- ▶ See [Beirlant *et al.*, 1997] for a review

Empirical estimates

► Pointwise estimates

Normalized histogram

$$p_{\text{emp}}^X(x) = \frac{1}{|\Omega|} \int_{\Omega} \delta(x - X(s)) ds$$

- **Biased** (even in low dimension)

Kernel density estimates (KDE)

- ▶ Count the # of samples within a given d -dimensional kernel centered at sample x :

Generalized KDE

$$p_{\text{KDE}}^X(x) = \frac{1}{|\Omega|} \int_{\Omega} K_{\sigma}(x - X(s)) \, ds$$

K_{σ} symmetric kernel (unit mass) with bandwidth
 $\sigma(\Omega, x, s)$

- Consistent but biased
- Bias depends on σ

Fixed bandwidth KDE

Parzen-Rosenblatt estimator

$$p_{\text{parzen}}^X(x) = K_\sigma * \frac{1}{|\Omega|} \int_{\Omega} \delta(x - X(s)) ds$$

\equiv smoothed normalized histogram

- Prototype: Gaussian kernel
- Bias **critically** depends on σ → **bandwidth estimation ?**
 - Plug-in estimator

$$\sigma = 0.9 \min \left(\hat{\sigma}, \frac{\hat{p}}{1.34} \right) |\Omega|^{-\frac{1}{5}}$$

→ **oversmoothing** when p^X has several modes

- $\hat{\sigma}$ empirical standard deviation of $X(\Omega)$
- \hat{p} interquartile range of $X(\Omega)$

- Cross-validation techniques
- On discrete square grids, sensitivity to **interpolation artefacts**
[Maes, 1998] [Pluim et al., 2000]

Fixed bandwidth KDE

► Regionalized KDE

- Improving estimation of minor modes of p^X which tend to be oversmoothed by global Parzen windowing

Localized Parzen-Rozenblatt estimator

$$p_{\text{parzen}}^X(x, s) = \frac{1}{|\Omega|} \int_{\Omega} W_{\sigma_s}(s' - s) K_{\sigma}(x - X(s')) ds'$$

K_{σ} symmetric kernel with bandwidth σ

W_{σ_s} symmetric spatial kernel with bandwidth σ_s
e.g. Gaussian kernel [Hermosillo, 2002] or

spatial window with finite extension σ_s
e.g. B-spline control patch [Loeckx et al., 2007]

- Bias dependence on σ is reduced but persists

Fixed bandwidth KDE

- ▶ Differentiable kernels yield differentiable density estimates
→ variational optimization of information-theoretic measures
- Non rigid registration: $X^\varphi = X \circ \varphi$ with transform $\varphi \in \mathcal{T}$
 - $\dim \mathcal{T} = \infty \rightarrow \varphi = \mathbb{I}d + \mathbf{u}$

$$\partial_{\mathbf{u}} p^{X^{\mathbf{u}}}(x) = \frac{1}{|\Omega|} \int_{\Omega} K'_\sigma(x - X(s)) \nabla X^{\mathbf{u}}(s) \, ds$$

[Hermosillo, 2002] [d'Agostino *et al.*, 2003] [Rougon *et al.*, 2005]

- $\dim \mathcal{T} < \infty \rightarrow \varphi(s; \Theta) = \mathbf{B}(s)\Theta$

$$\partial_{\Theta} p^{X^\varphi}(x) = \frac{1}{|\Omega|} \int_{\Omega} K'_\sigma(x - X(s)) [\nabla X^\varphi(s)^T \mathbf{B}(s)] \, ds$$

[Viola, 1995] [Rueckert *et al.*, 1999]

Fixed bandwidth KDE

- ▶ Differentiable kernels are **not** required for the variational optimization of information-theoretic measures over shape spaces

- Active region segmentation: X^Γ with region boundary $\Gamma \in \mathcal{S}$
Level set framework: $\Gamma = \{ s \in \Omega \mid \phi(s) = 0 \} \rightarrow X^\phi$
- $\dim \mathcal{S} = \infty$

$$\partial_\phi p^{X^\phi}(x) = \frac{1}{|\Omega|} \left(p^{X^\phi}(x) - K_\sigma(x - X^\phi(s)) \right) \delta(\phi(s))$$

[Freedman, 2004] [Kim *et al.*, 2005] [Herbulot *et al.*, 2006]
[Rougon *et al.*, 2006] [Michailovich *et al.*, 2007]

- $\dim \mathcal{S} < \infty \rightarrow \phi(s; \Theta) = B(s)\Theta$

$$\partial_\Theta p^{X^\phi}(x) = \frac{1}{|\Omega|} \left(p^{X^\phi}(x) - K_\sigma(x - X^\phi(s)) \right) B(s)^T \delta(\phi(s))$$

[Bernard *et al.*, 2009]

Fixed bandwidth KDE

► Curse of dimensionality: as $d \nearrow$

- the sample set $X(\Omega)$ gets sparser in the state space \mathcal{X}
- $|\Omega|$ must \nearrow exponentially to ensure that enough samples fall within the kernel
- if $|\Omega|$ is fixed, the performance of p_{parzen}^X rapidly degrades
[Cacoullos, 1966] [Epanechnikov, 1969]

Parzen density estimator is unreliable for $d > 3$

Adaptive bandwidth KDE

► Bandwidth adaptation strategies [Terrell and Scott, 1992]

- Sample-dependent: $\sigma(X_s) \rightarrow$ sample point estimators
 - $\sigma(X_s) \propto p^X(X_s)^{-1/2}$ [Abramson, 1982]
 - See also [Comaniciu, 2003]
- Estimation point-dependent: $\sigma(x) \rightarrow$ balloon estimators

Direct approach

- adapt bandwidth around x to include enough samples
- see for instance [Sain, 2002]

Dual approach

- set the # of samples per volume unit to a given value k
- compute the containing ball $B_k^X(x)$

→ kNN density estimators

Foundations > Bibliography |

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Foundations > Bibliography IV

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Outline

- 1 Foundations
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- 3 Entropic graphs
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kth nearest neighbor (kNN) density estimate

kNN density estimator [Loftsgaarden and Quesenberry, 1965]

$$p_{\text{knn}}^X(x) = \frac{k}{|\Omega| V_d (\rho_k^X(x))^d}$$

$\rho_k^X(x)$ Euclidean distance from x to its k th nearest neighbor in the sample set $X(\Omega) \setminus \{x\}$

V_d Volume of the unit ball of \mathbb{R}^d : $V_d = \frac{\pi^{d/2}}{\Gamma(d/2+1)}$

- $B_k^X(x)$: d -dimensional ball of radius $\rho_k^X(x)$ centered at x
- Consistent but biased (even in low dimension)
- Bias depends on k

→ Not used for density estimation but yields consistent and asymptotically unbiased estimators of entropy in high dimensions

kNN vs. KDE

Relationship between KDE and kNN density estimates

$p_{\text{knn}}^X(x)$ can be interpreted as a **balloon estimate** with **uniform kernel** $U_{\rho_k^X(x)}$ over the ball $\mathcal{B}_k^X(x)$ with bandwidth $\sigma(x) = \rho_k^X(x)$

From kNN density to kNN entropy estimation

- ① Consider the Ahmad-Lin entropy estimator

$$H^{\text{AL}}(X) = -\frac{1}{|\Omega|} \sum_{s \in \Omega} \log p^X(X_s)$$

- ② Plug p_{knn}^X . This yields a geometric entropy estimator:

$$H^{\text{knn}}(X) = \frac{1}{|\Omega|} \sum_{s \in \Omega} \log \left[\rho_k^X(X_s) \right]^d + \log \frac{V_d |\Omega|}{k}$$

- Consistent but biased
- Bias depends on k
- Bias \searrow as $d \nearrow$
- H^{knn} performs significantly better than fixed bandwidth KDE as $d \nearrow$

- ③ Fix the bias

kNN entropy estimators

Kosachenko-Leonenko-Goria entropy estimator [Goria *et al.*, 2005]

Under weak conditions on X , the following estimator is a **consistent** and **asymptotically unbiased** estimator of differential entropy:

$$H^{\text{knn}}(X) = \frac{1}{|\Omega|} \sum_{s \in \Omega} \log \left[\rho_k^X(X_s) \right]^d + c_k(d)$$

where $c_k(d) = \log(V_d(|\Omega| - 1)) - \psi(k)$ and $\psi(k) = \frac{\Gamma'(k)}{\Gamma(k)}$

- Conditions on k :

- $\lim_{|\Omega| \rightarrow \infty} k = \infty$ and $\lim_{|\Omega| \rightarrow \infty} \frac{k}{|\Omega|} = 0$ e.g. $k = \sqrt{|\Omega|}$
- Image analysis: $k > \#$ degrees of freedom

- Low sensitivity to k

f -entropy kNN estimators

- Tsallis / Renyi entropies involve the integral:

$$\mathcal{I}_q^X = \int (p^X(x))^q dx = \mathbb{E}_X \left[(p^X)^{q-1} \right]$$

Lemma [Leonenko *et al.*, 2008]

Under weak conditions on X , the following estimator is a **consistent** and **asymptotically unbiased** kNN estimator of \mathcal{I}_q^X :

$$\widehat{\mathcal{I}_q^X} = \frac{1}{|\Omega|} \sum_{s \in \Omega} c_k(d) \left([\rho_k^X(X_s)]^d \right)^{1-q}$$

where $c_k(d) = (V_d(|\Omega| - 1))^{1-q} \frac{\Gamma(k)}{\Gamma(k+1-q)}$

f -entropy kNN estimators

Tsallis / Renyi entropy kNN estimator [Leonenko *et al.*, 2008]

The following estimators are **consistent** and **asymptotically unbiased** estimators of Tsallis / Renyi entropies, respectively:

$$H_{\alpha}^{\text{knn}}(X) = \frac{1}{\alpha - 1} \left(1 - \widehat{\mathcal{I}}_{\alpha}^X \right)$$

$$H_r^{\text{knn}}(X) = \frac{1}{1 - r} \log \widehat{\mathcal{I}}_r^X$$

KL divergence kNN estimators

KL divergence kNN estimator [Goria *et al.*, 2005]

Under weak conditions on X , the following estimator is a **consistent** and **asymptotically unbiased** estimator of KL divergence $D_{\text{KL}}(X)$:

$$D_{\text{KL}}^{\text{knn}}(X \parallel Y) = \frac{1}{|\Omega|} \sum_{s \in \Omega} \log \left[\frac{\rho_k^Y(X_s)}{\rho_k^X(X_s)} \right]^d + c_k(d)$$

where $c_k(d) = \log \frac{|\Omega|}{|\Omega|-1}$

Note: This estimator involves the d -dimensional ball $\mathcal{B}_k^Y(X_s)$ comprising the k nearest samples of X_s in the sample set $Y(\Omega)$

- See also [Perez-Cruz, 2008] [Wang *et al.*, 2009]

f -divergence kNN estimators

- Chernoffs / Renyi divergences involve the integral:

$$\mathcal{I}_q^{X||Y} = \int_{\Omega} [p^X(x)]^q [p^Y(x)]^{1-q} dx = \mathbb{E}_X \left[\left(\frac{p^X}{p^Y} \right)^{q-1} \right]$$

Lemma [Leonenko et al., 2008]

Under weak conditions on X , the following estimator is a **consistent** and **asymptotically unbiased** kNN estimator of $\mathcal{I}_q^{X||Y}$:

$$\widehat{\mathcal{I}_q^{X||Y}} = \frac{1}{|\Omega|} \sum_{s \in \Omega} d_k(d) \left(\left[\frac{\rho_k^Y(X_s)}{\rho_k^X(X_s)} \right]^d \right)^{1-q}$$

$$\text{where } d_k(d) = (V_d |\Omega|)^{1-q} \frac{\Gamma(k)}{\Gamma(k+1-q)}$$

f -divergence kNN estimators

Chernoff / Renyi divergence kNN estimator [Goria *et al.*, 2005]

The following estimators are **consistent** and **asymptotically unbiased** estimators of Chernoff / Renyi divergences, respectively:

$$D_{\alpha}^{\text{knn}}(X \parallel Y) = \frac{1}{\alpha - 1} \left(1 - \widehat{\mathcal{I}_{\alpha}^{X \parallel Y}} \right)$$

$$D_r^{\text{knn}}(X \parallel Y) = \frac{1}{1-r} \log \widehat{\mathcal{I}_r^{X \parallel Y}}$$

Mutual information kNN estimators

Mutual information kNN estimator [Hamrouni and Rougon, 2010]

Under weak conditions on X , the following estimator is a **consistent** and **asymptotically unbiased** estimator of mutual information $I(X)$:

$$I^{\text{knn}}(X, Y) = \frac{1}{|\Omega|} \sum_{s \in \Omega} \log \left[\frac{\rho_k^X(X_s) \rho_k^Y(Y_s)}{(\rho_k^{X,Y}(X_s), Y_s))^2} \right]^d + c_k(d)$$

$$\text{where } c_k(d) = \log \left(\frac{V_d^2}{V_{2d}} (|\Omega| - 1) \right) - \psi(k)$$

Note: I^{knn} involves the $2d$ -dimensional ball $\mathcal{B}_k^{X,Y}(X_s, Y_s)$ containing the k nearest samples of (X_s, Y_s) in the sample set $X(\Omega) \times Y(\Omega)$

- Alternative kNN estimators: [Kraskov *et al.*, 2004] [Evans, 2007]

Optimizing kNN entropy estimators

- ▶ kNN estimators are **not differentiable**
 - no direct variational optimization
- ▶ A **plug-in approach** is used [Boltz *et al.*, 2009]
 - ➊ Compute the criterion derivative classically using differentiable Parzen estimators K_σ
 - ➋ Replace K_σ by uniform kernel $U_{\rho_k^X(x)}$ over kNN ball $\mathcal{B}_k^X(x)$
 - integration over K_σ support → **finite sum** over $\mathcal{B}_k^X(x)$
 - K_σ derivative → indicator $1_{\mathcal{S}_k^X(x)}$ along **kNN sphere** $\mathcal{S}_k^X(x)$

Optimizing kNN entropy estimators

- Non rigid registration: $X^\varphi = X \circ \varphi$

Let $x = X^\varphi(s_0)$ and $y = X^\varphi(s)$

$$\partial_\varphi p^{X^\varphi}(x) = \frac{1}{|\Omega|} \int_{\Omega} K'_\sigma(x - y) \nabla X^\varphi(s) ds$$

translates into:

$$\partial_\varphi p^{X^\varphi}(x) = \frac{1}{|\Omega|} \sum_{s \in \Omega} U'_{\rho_k^X(x)}(x - y) \nabla (X^\varphi(s_0) - X^\varphi(s))$$

with:

$$U'_{\rho_k^X(x)}(x - y) = \frac{-1}{V_d(\rho_k^X(x))^d} \frac{x - y}{|x - y|} \mathbf{1}_{S_k^X(x)}(y)$$

Note: At most k points lie on the kNN sphere $S_k^X(x)$

→ low computational cost

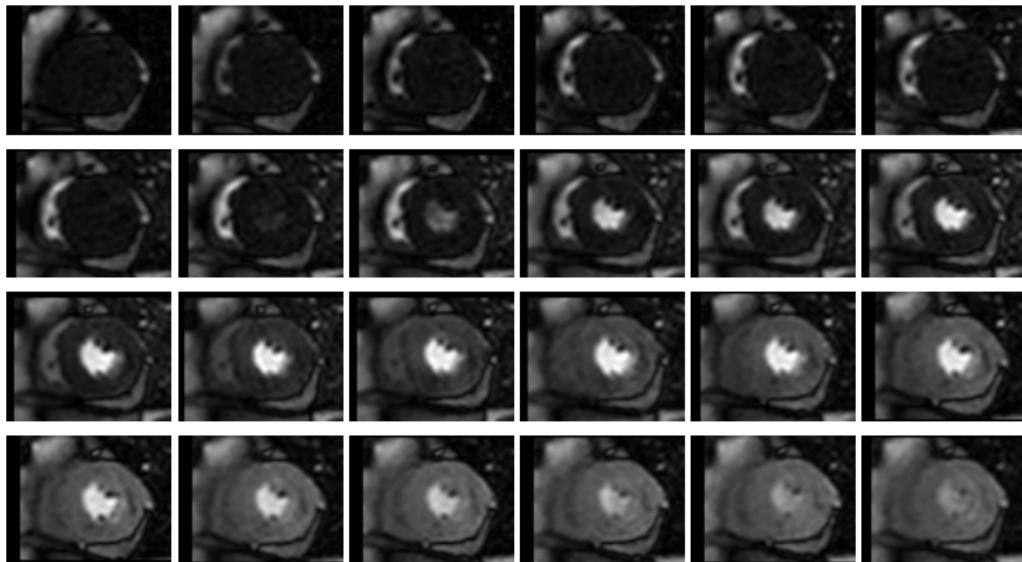
Efficient kNN search

► Computational issues

- The numerical complexity of kNN-based estimation and optimization is dictated by the kNN search algorithm
- Exact kNN search (kd-tree based) can be prohibitive
- Approximate Nearest Neighbors (ANN) algorithm
[Arya *et al.*, 1998]
 - $o(|\Omega|)$ and $o(d)$ complexity and memory usage
 - Free C++ ANN library: www.cs.umd.edu/~mount/ANN
 - GPU implementation in CUDA
[Garcia *et al.*, 2008]

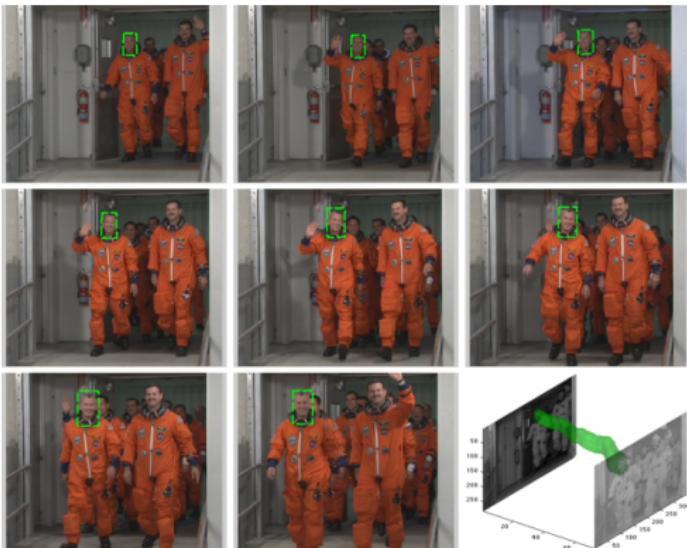
Example applications

- Groupwise rigid registration using MI kNN estimator
[Hamrouni and Rougon *et al.*, 2010]



Example applications

- Tracking using KL divergence kNN estimator
[Boltz *et al.*, 2009]



Source: E. Debreuve (www.i3s.unice.fr/~debreuve)

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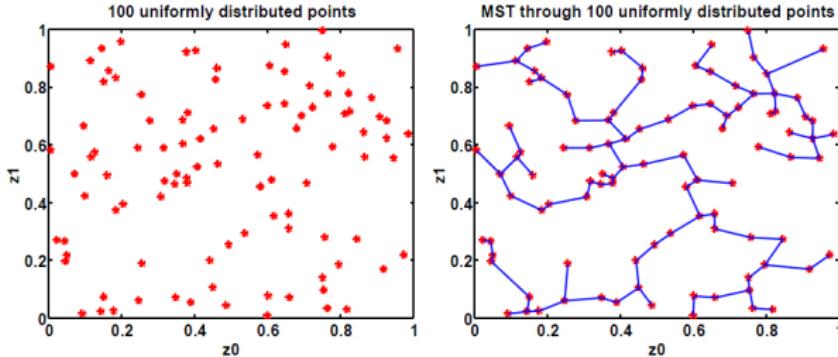
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- 4 Multivariate information measures

Spanning graphs

► Describing sample sets as spanning graphs

[Readmond and Yukich, 1996] [Steele, 1997] [Yukich, 1998]

- The sample set $X(\Omega)$ is viewed as the set of d -dimensional vertices $V = (X_s)_{s \in \Omega}$ of a graph $G = (V, E)$ with edges E
- The graph set \mathcal{G} is the set of **spanning graphs** (i.e. connected, acyclic) with vertices V



From [Neemuchwala and Hero, 2005]

Minimal spanning graphs

- Length of a spanning graph $G \in \mathcal{G}$:

$$\mathcal{L}_\gamma(G) = \sum_{e \in E} \|e\|^\gamma$$

$\|e\|$ Euclidean length of vertex e

γ power weighting constant $(0 < \gamma < d)$

- Minimal spanning graphs: graphs minimizing $\mathcal{L}_\gamma(G)$
Denote by $L_\gamma(G)$ the minimal length:

$$L_\gamma(G) = \min_{G \in \mathcal{G}} \mathcal{L}_\gamma(G)$$

- Computing $L_\gamma(G)$ requires a combinatorial optimization over the set of spanning graphs \mathcal{G}

Entropic graphs

Theorem [Beardwood-Halton-Hammersley, 1959]

If $X(\Omega)$ is a set of i.i.d. samples drawn from a d -dimensional density p^X and $L_\gamma(G)$ is continuous quasi-additive, then:

$$\lim_{|\Omega| \rightarrow \infty} \frac{L_\gamma(G)}{|\Omega|^r} = \beta_{d,\gamma} \int (p^X(x))^r dx \quad (\text{a.s.})$$

where $r = \frac{d-\gamma}{d}$ and $\beta_{d,\gamma}$ is a constant independent of p^X

Entropic graphs [Hero and Michel, 1999]

$$H_r^{\text{ent}}(X) = \frac{1}{1-r} \log \left(\frac{L_\gamma(G)}{|\Omega|^r} \right) - c$$

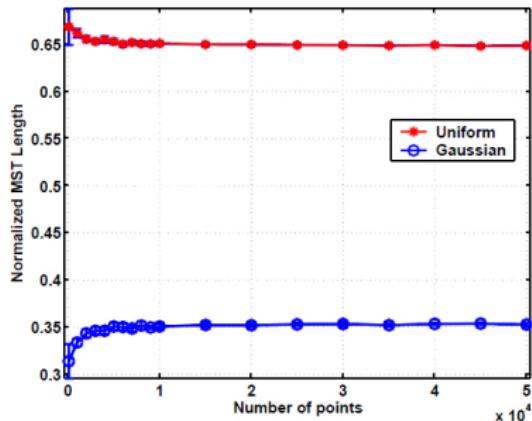
is a consistent and asymptotically unbiased estimator of Renyi entropy $H_r(X)$, where $c = \frac{1}{1-r} \log \beta_{d,\gamma}$ is a bias correction



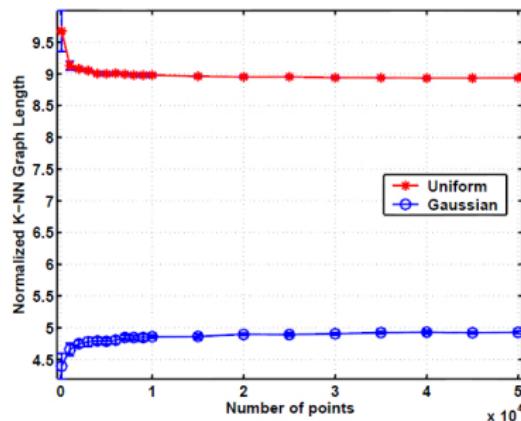
Computing entropic graphs

- $H_r(X)$ can be estimated from $L_\gamma(G)$ for **any** minimal spanning graph with continuous quasi-additive length functional

MST graph



kNN graph



From [Neemuchwala and Hero, 2005]

Computing entropic graphs

- Focus on efficiently computable graphs
 - Minimal Spanning Tree → kNN Kruskal algo.: $o(|E| \log |E|)$
[Hero and Michel, 1999] [Neemuchwala *et al.*, 2005b]
 - kNN graph → ANN algorithm: $o(|\Omega|)$
[Neemuchwala *et al.*, 2005]
- Once G is computed, $H_r(X)$ can be estimated for different values of $r \in [0, 1]$ by adjusting $\gamma = d(1 - r)$
- The topology of G is independent of γ in many cases
→ Single optimization required to estimate $H_r(X)$ for all r
- The bias correction c (which depends on graph type) can be estimated offline by Monte-Carlo simulation

Entropic graph estimators

- Entropic graphs yield **geometric parameter-free** estimators of:

- Renyi divergence D_r and information I_r

Recall: $D_1 = \text{KL}$ divergence, $I_1 = \text{mutual information}$

- Generalized Jensen divergence

$$\Delta_{r,p}(X, Y) = H_r(pX + qY) - [pH_r(X) + qH_r(Y)]$$

- Generalized geometric-arithmetic mean divergence

$$D_{r,p}^{\text{GA}}(X \parallel Y) = D_r(pX + qY \parallel X^p Y^q)$$

where $p \in [0, 1]$ and $q = 1 - p$

Entropic graph estimators

Entropic graph estimator of Renyi information

$$I_r^{\text{ent}}(X, Y) = \frac{1}{r-1} \log \frac{1}{|\Omega|^r} \sum_{s \in \Omega} \sum_{j=1}^k \left(\frac{\rho_j^{X,Y}(X_s, Y_s)}{\rho_j^X(X_s) \rho_j^Y(Y_s)} \right)^{2\gamma}$$

is a **consistent** and **asymptotically unbiased** estimator of Renyi information $I_r(X, Y)$

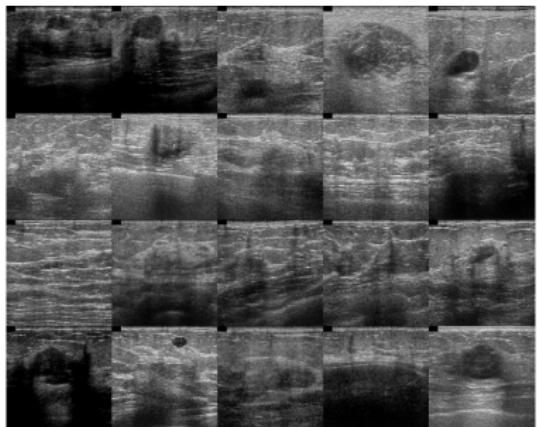
Note: $I_r^{\text{ent}}(X, Y)$ involve on 3 minimal graphs, with respective vertices $(X_s, Y_s) \in \mathbb{R}^{2d}$, $X_s \in \mathbb{R}^d$ and $Y_s \in \mathbb{R}^d$

Optimizing entropic graph estimators

- ▶ Non rigid registration: $X^\varphi = X \circ \varphi$ with $\varphi(s; \Theta) = \mathbf{B}(s)\Theta$
 - Closed-form expression of the derivative of kNN entropic graph estimators w.r.t. Θ
 - Renyi joint-entropy [Oubel *et al.*, 2007]
 - Renyi information [Staring *et al.*, 2009]
 - Iterative variational optimization
 - Assumption: **invariance of graph topology** for small $\delta\Theta$
 - See also [Sabuncu and Ramadge, 2008]
 - Efficient stochastic gradient scheme [Staring *et al.*, 2009]
 - Much faster than finite difference or direct optimization
 - Yet **computationally intensive**

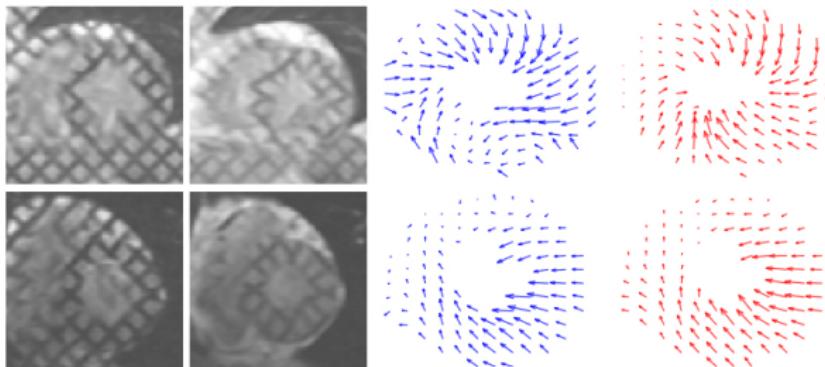
Example applications

- ▶ Breast ultrasound image registration
[Neemuchwala *et al.*, 2007]
- Features: ICA coefficients ($d = 64$)
Alternatives: wavelet coefficients, curvelet coefficients, image patches
- Transforms: isometries
- Generalized Jensen divergence
- kNN entropic graphs



Examples applications

- ▶ Cardiac tagged MR image registration
[Oubel *et al.*, 2005] [Oubel *et al.*, 2006]
- Features: one-level Haar DWT ($d = 4$) or CWT ($d = 6$)
- Transforms: multilevel FFD
- Renyi information
- kNN entropic graphs

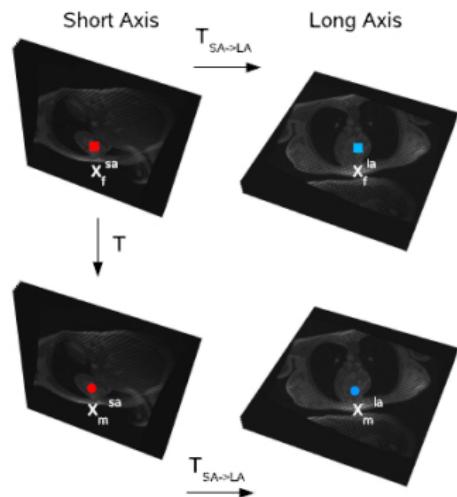


kNN entropic graph L_r vs. standard NMI

Examples applications

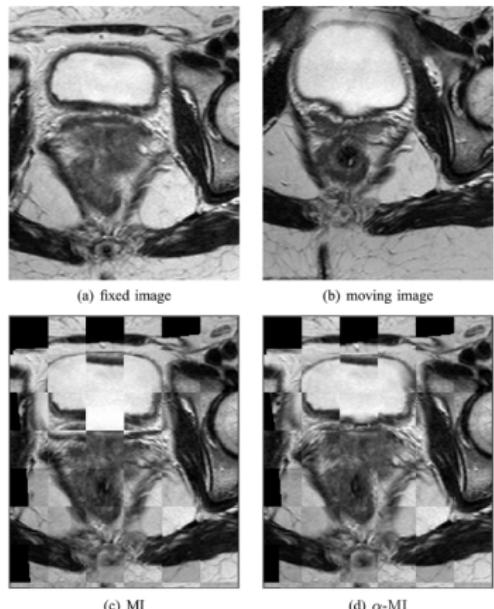
- ▶ Multiview cardiac tagged MR image registration
[Oubel *et al.*, 2007]

- Features: graylevel in SA and LA views ($d = 2$)
- Transforms: multilevel FFD
- Renyi joint entropy
- kNN entropic graphs
- Variational optimization



Examples applications

- ▶ Cervical MRI registration
[Staring *et al.*, 2009]
- Features: intensity + 2nd-order 3D local image invariants at 2 scale levels ($d = 15$). PCA on image union $\rightarrow d = 6$
- Transforms: FFD
- Renyi information
- kNN entropic graphs
- Variational optimization



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Outline

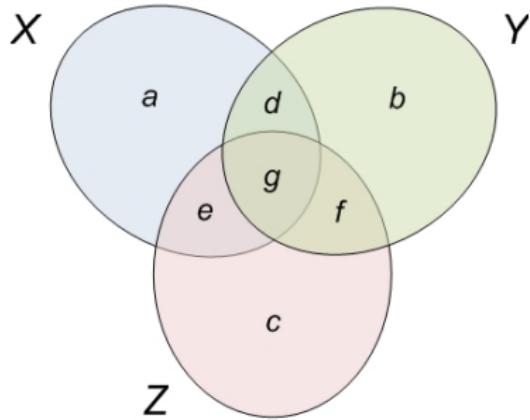
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Multivariate information measures

- ▶ Multivariate information-theoretic measures are appealing for multiple image analysis (e.g. multi-view, multi-frame...)
 - groupwise registration
 - registration using class information
 - spatio-temporal segmentation
 - ...

But generalizing bivariate information measures to the multivariate case is **not straightforward**

Co-information



Co-information is the information shared by **all** RVs

→ g

$$\begin{aligned} I(X, Y|Z) - I(X, Y) &= -H(X) - H(Y) - H(Z) \\ &\quad + H(X, Y) + H(X, Z) + H(Y, Z) \\ &\quad - H(X, Y, Z) \end{aligned}$$

Co-information

Denote $\mathcal{X} = \{X^1, X^2, \dots, X^n\}$ and $\mathcal{Y} \subseteq \mathcal{X}$ a RV subset of \mathcal{X}

Co-information [McGill, 1954]

$$\begin{aligned} I(X^1, \dots, X^n) &= \sum_{\mathcal{Y} \subseteq \mathcal{X}} (-1)^{|\mathcal{X} \setminus \mathcal{Y}|} H(\mathcal{Y}) \\ &= I(\mathcal{X} \setminus X^i | X^i) - I(\mathcal{X} \setminus X^i) \quad \forall X^i \in \mathcal{X} \end{aligned}$$

- Also called **information interaction**
[Bell, 2003] [Kludas *et al.*, 2009] [Zhou and Li, 2010]
- Symmetric
- Stable and **unambiguous**: adding new RVs does not change existing interactions (can only add new ones)

Co-information

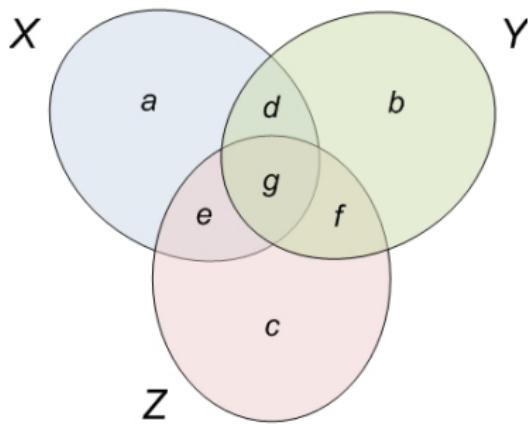
Denote $\mathcal{X} = \{X^1, X^2, \dots, X^n\}$ and $\mathcal{Y} \subseteq \mathcal{X}$ a RV subset of \mathcal{X}

Co-information [McGill, 1954]

$$\begin{aligned} I(X^1, \dots, X^n) &= \sum_{\mathcal{Y} \subseteq \mathcal{X}} (-1)^{|\mathcal{X} \setminus \mathcal{Y}|} H(\mathcal{Y}) \\ &= I(\mathcal{X} \setminus X^i | X^i) - I(\mathcal{X} \setminus X^i) \quad \forall X^i \in \mathcal{X} \end{aligned}$$

- Can be positive (synergy) or negative (redundancy)
- Involve all p -dimensional ($p \leq d$) densities $p^{\mathcal{Y}}$ ($\mathcal{Y} \subseteq \mathcal{X}$)
→ high estimation / optimization computational cost
- Often, too stringent similarity constraint
- Currently unused in medical imaging

Multi-information



Multi-information is the information shared by at least 2 RVs

$$\rightarrow \sum -a - b - c$$

$$H(X) + H(Y) + H(Z) - H(X, Y, Z)$$

Multi-information

Non rigid registration [Studholme *et al.*, 2006]

→ measurement of volume change in brain MRI

- X, Y : images
 Z : (overlapping) region labels
- Assumption: (X, Y) independent of Z
≡ regionalized MI
- Regionalized KDE
- Transform: fluid
- Variational optimization

Multi-information

In the same spirit:

- Multi-channel MI [Holden *et al.*, 2004]

$$MI(X, X', Y, Y') = H(X, X') + H(X, Y') - H(X, X', Y, Y')$$

Note: X' (Y') are 1st- or 2nd-order scale-space derivatives of X (Y)

- Trivariate NMI [Papp *et al.*, 2009]

$$NMI(X, Y, Z) = \frac{H(X) + H(Y) + H(Z)}{H(X, Y, Z)}$$

Multi-information

Multi-information [Studeny and Vejnarova, 1998]

$$\begin{aligned} I(X^1, \dots, X^n) &= \sum_i H(X^i) - H(X^1, \dots, X^n) \\ &= \int p^{X^1, \dots, X^n} \log \frac{p^{X^1, \dots, X^n}}{p^{X^1} \dots p^{X^n}} d\mathbf{x} \end{aligned}$$

- Also called **total correlation**
- Symmetric, nonnegative
- Redundant information overcount
- Involves ***nd*-dimensional** joint density
- Straightforward generalization to arbitrary information metrics

Generalized multi-information

Integral generalized multi-information

$$I_f(X^1, \dots, X^n) = \int p^{X^1} \dots p^{X^n} f \left(\frac{p^{X^1, \dots, X^n}}{p^{X^1} \dots p^{X^n}} \right) d\mathbf{x}$$

f continuous, convex over \mathbb{R}^+

Non-integral generalized multi-information

$$I_\psi(X^1, \dots, X^n) = \log \psi^{-1} \left(\int p^{X^1, \dots, X^n} \psi \left(\frac{p^{X^1, \dots, X^n}}{p^{X^1} \dots p^{X^n}} \right) d\mathbf{x} \right)$$

ψ continuous, monotonic over \mathbb{R}^+

Generalized multi-information estimators

- Consistent and asymptotically unbiased kNN entropy estimators / entropy graphs estimators are easily derived

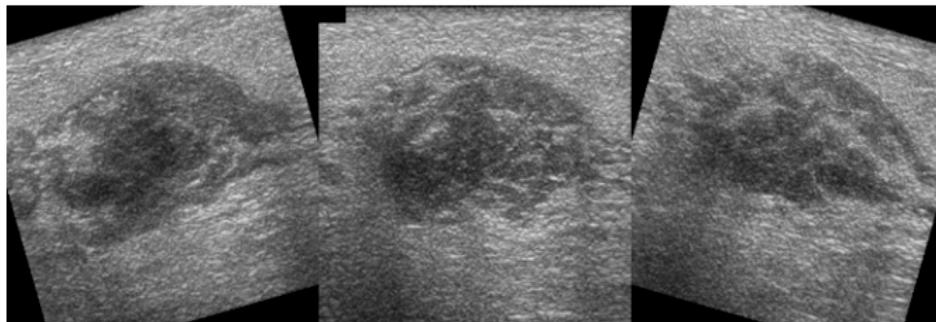
Entropic graph estimator of Renyi multi-information
[Neemuchwala *et al.*, 2007]

$$I_r^{\text{ent}}(X^1, \dots, X^n) = \frac{1}{r-1} \log \frac{1}{|\Omega|^r} \sum_{s \in \Omega} \sum_{j=1}^k \left(\frac{\rho_j^{X^1, \dots, X^n}(x_s^1, \dots, x_s^n)}{\rho_j^{X^1}(x_s^1) \dots \rho_j^{X^n}(x_s^n)} \right)^{2\gamma}$$

Note: Involves $n + 1$ minimal graphs

Example application

- ▶ Breast ultrasound multiple image registration
[Neemuchwala *et al.*, 2007]



- Features: ICA coefficients ($d = 64$)
- Transforms: isometries
- Renyi multi-information
- kNN entropic graphs

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Medical image analysis using high-dimensional information-theoretic criteria

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